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Affirmative Action Policies for Top-k Candidates Selection

With an Application to the Design of Policies for University Admissions

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ABSTRACT

We consider the problem of designing affirmative action policies for selecting the top- k candidates from a pool of applicants. We assume that for each candidate we have socio-demographic attributes and a series of variables that serve as indicators of future performance (e.g., results on standardized tests) – as well as historical data including the actual performance of previously selected candidates. We consider the case where an organization wishes to increase the selection of people from disadvantaged socio-demographic groups. Hence, we seek to design an *affirmative action* policy to select candidates who are more likely to perform well, but in a way that increases the representation of disadvantaged groups.

Our motivating application is the design of university admission policies to bachelor's degrees. We use a causal framework to describe several families of policies (changing component weights, giving bonuses, and enacting quotas), and compare them both theoretically and through extensive experimentation on a real-world dataset containing thousands of university applicants. Our empirical results indicate that simple policies could favor the admission of disadvantaged groups without significantly compromising on the quality of accepted candidates.

CCS CONCEPTS

- **Computing methodologies** → *Model development and analysis*;
- **Applied computing** → *Computers in other domains*;

KEYWORDS

algorithmic fairness, data-driven policy design

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1 INTRODUCTION

The emergent academic field of algorithmic fairness is concerned with the impact of algorithmic decision making on society, particularly over disadvantaged groups.

Predictive policies for top-k selection. We consider a general setting in which we have a large pool of applicants for, e.g., an educational program, a scholarship, or a job, and we want to select the k most promising ones. This selection is based on the candidates' predicted future performance. For instance, in the case of university admissions, this prediction stems from previous grades obtained in different subjects (transcripts), scores in relevant standardized tests, ratings from interviews with the candidate, and so on. The predictor, which induces a *selection policy*, is created from historical data about the academic performance of previously selected candidates. A key constraint is that such outcomes are only observed for candidates that have been selected in the past – but not for rejected ones.

Affirmative action in top-k selection. Algorithmic fairness is commonly operationalized through a notion of *demographic parity* or *statistical parity*, which requires that a decision such as accepting or rejecting a candidate is independent of the sensitive attribute (for a criticism of this notion see [7]). In our setting, for example, we want to admit candidates of different groups at similar rates. Given that disparities exist, we seek an *affirmative action* policy: a temporary intervention aimed at increasing the representation of an under-represented group.

University admissions. The problem domain we use throughout this paper is university admission policies. In several countries, standardized tests for university admission are administered at the end of the last year of high school or community college. They are often framed as a measurement of aptitude for university studies, i.e., a prediction of how a student would eventually perform if admitted. For most students, admission decisions are based on an *admission score* computed from test scores and/or high school grades (e.g., through a linear combination).

Admission scores offer one mechanism for admission targeted at aptitude. However, universities may also want to have a diverse body of students, or may wish to promote gender equality and social mobility. For this reason, university admissions often include a predictor of students' future performance plus a patchwork of additional affirmative action programs addressed at increasing the number of accepted candidates from disadvantaged backgrounds.

Designing affirmative action policies. The goals of an affirmative action top-k selection policy are twofold. First, it should select candidates who have a high expected performance. Second, it should ensure a sufficient number of candidates from disadvantaged

backgrounds are selected. Both properties can be seen as expressing notions of fairness: the former towards good performance; the latter against disadvantages that are not related with performance.

Effectively, to design such a policy one needs to predict its effect before it is enacted. In this paper, we formalize the problem of designing data-driven affirmative action policies and to develop algorithms to build such policies starting from historical data. Our contributions have direct application to university admissions, but they are directly extensible to other domains.

Our approach. To formalize the questions we consider, we build a probabilistic causal model that captures the selection process. The causal model takes into account each applicant’s input attributes (e.g., test scores), as well as a number of *sensitive* socio-demographic attributes, to compute an admission score, which is the main basis for their selection. Given the causal model, we consider three types of policies: (1) **COEFFICIENTS**: policies that adjust the weight of different components of the admission score to emphasize those that favor disadvantaged groups; (2) **BONUS**: policies that give extra points in the admission score to applicants from disadvantaged groups; and (3) **QUOTA**: policies that enforce a certain selection rate for applicants from disadvantaged groups. Policies are considered from the perspective of a single institution and optimized to increase the expected performance of the selected applicants while accepting a suitable proportion of disadvantaged applicants.

In our analysis, we build instances of causal models from the available data, find suitable parameters for a series of affirmative action policies, and present the results of an empirical study over university admissions considering thousands of applicants.

2 RELATED WORK

2.1 Fairness-aware top-k selection

Ranking with fairness constraints, in batch processing and online processing scenarios respectively, is the focus of two recent works by Celis et al. [4] and Stoyanovich et al. [22]. In these works, every element to be ranked has a certain utility and belongs to a certain group, and for every ranking position $r = 1, 2, \dots, k$ there are two constraints per group indicating the minimum and maximum number of candidates from that group that must be among the top- r positions. The goal is to rank candidates by decreasing utility without violating the constraints.

Zehlike et al. [24] define a *fair representation condition* based on a binomial test that seeks to determine if a ranking of k elements is compatible with a process in which candidates have been drawn at random from protected and non-protected groups with a known probability. The fair representation condition together with a process for adjusting for multiple hypothesis testing can be used to obtain the bounds used in ranking with fairness constraints. Kearns et al. [14] study the case of selecting k candidates drawn from d different groups in which candidates across groups are not comparable. Hence, what matters is the relative position of candidates within their groups.

2.2 Affirmative action

“Downstream” effects of affirmative action. Several recent papers study “downstream” or “cascade” effects of affirmative action

policies, mostly through a theoretical approach. Hu and Chen [12] consider a labor market in which firms are either temporary or permanent employers. Temporary employers are forced to use hiring policies that favor disadvantaged groups, while permanent employers can hire based on expected utility alone. The goal is to increase the employability of candidates from disadvantaged groups through their employment in the temporary market. Kannan et al. [13] analyze the potential effects of affirmative action in education on the employability of graduates from disadvantaged groups. The goal is to achieve parity of employability or create an incentive for employers to adopt a color-blind hiring policy. Mouzannar et al. [18] analyze the effect of affirmative action policies on the qualifications of different groups in society.

Data-driven analysis. Over the last four decades prestigious universities, many of which are public institutions, have become even more selective, having a negative effect on students from low socioeconomic backgrounds and underrepresented minorities [20]. This effect is mainly explained by the lower performance of disadvantaged students on standardized tests and other admission requirements that favor applicants from well-off families [3]. However, admission scores might underestimate university performance for disadvantaged students. For instance, a study by Wightman [23] on nearly 25,000 law students in the US found that while students of color had significantly poorer results on admission tests than white students, their difference in terms of performance once they were admitted (their probability of passing the bar exam) did not justify the large gap between the admission rates of both groups.

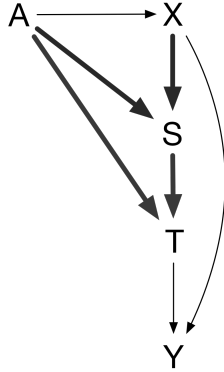
As a result, countries and universities have promoted affirmative action policies. For instance, Brazil has promoted bonus [8] and racial quotas [10]. The evidence suggests that the implementation of these policies increased both black students and students coming from low socioeconomic backgrounds in large public universities [10]. In Chile, where significant gaps are observed in standardized math tests between students of public versus private schools [5] and between women and men [1], coefficient-based policies have also shown some effectiveness [17].

2.3 Causal reasoning and fairness

Hardt et al. [11] explain how causal reasoning, as developed by Pearl [19] and others, offers a theoretical framework to express notions of algorithmic fairness that go beyond statistical / observational measures. Indeed, the work of Kusner et al. [16] and Kilbertus et al. [15] builds upon that intuition by studying notions of *counterfactual fairness* that take into account the causal structure of the data’s generative model. In our work, we use a causal framework to define the model for the admission mechanism (Section 3).

3 SETTING

In the setting we consider, selection of candidates is performed by *one institution*. We begin by defining a probabilistic causal model that captures this process, expresses our assumptions about the data, and allows us to quantify the effect of different selection policies (Section 3.1). For a given policy, we then define measures of *utility*, capturing the performance of selected candidates, and *fairness*, capturing the similarity between selection rates (Section 3.2). Next, we describe three types of policies (Section 3.3), and present a



Symbol	Description
A	Socio-economic attributes
X	Input scores (e.g., results of qualification tests)
S	Selection score: combination of input scores
T	Decision of the institution: select or not select
Y	Performance of selected candidate

Figure 1: Diagram of causal model. Selection policies are defined via the causal relationships $(A, X) \rightarrow S$ and $(A, S) \rightarrow T$. Relationships affected by the affirmative action policy are shown with thick arrows.

formal definition of the technical problem we address as a trade-off between utility and fairness (Section 3.4). Finally, we discuss various modeling choices we make (Section 3.5).

3.1 Causal model

Our design of selection policies is guided by historical data that are the result of a particular enacted selection policy. To predict what data we would observe under alternative policies, it is necessary to express formally our knowledge and assumptions about the decision mechanisms we study. We do this by defining a *causal diagram* [19] capturing our background knowledge about causal relationships between quantities of interest, as shown in Figure 1. The diagram describes the selection mechanism from the point of view of *one candidate* to *one institution*, and assumes that each candidate is associated with socio-economic attributes A , which are considered sensitive, and a series of input scores X (e.g., results of qualification tests), which are considered non-sensitive. These are combined into a selection score S which is the basis for the institution’s decision T . In particular, given input scores $X = x$ and socio-economic attributes $A = a$, the candidate obtains a selection score $S = s$ according to $S = f(X, A)$. Here, we allow the selection score S to depend on A to allow for affirmative action through the selection score function. Based on this score S a decision T is made in favor or against the candidate. Again, we allow T to be determined by A so as to allow also for affirmative action through policies that operate at the stage of decision.

For the purposes of affirmative action policy design, it is also crucial that we model the possibility that sensitive socio-economic attributes A influence the input scores X . This is done through

the dependency $A \rightarrow X$ in the diagram. This is the case if, for instance, university applicants from families of higher income have access to training or resources that are not affordable for families of lower income. In that case, an affirmative action policy tries to bring *equality of opportunity* [21].

The institution under consideration is interested in the acceptance of only a certain number or fraction of top candidates. Formally, we require every selection policy to be *calibrated* – i.e., that it lead, in expectation, to the selection of a certain fraction θ of candidates, considered known and fixed for the institution.

$$\mathbb{E}[T] = P(T = 1) = \theta \quad (1)$$

If the decision for a candidate is positive ($T = 1$), we observe their performance Y . If a candidate is rejected the performance of that candidate is not observed, and we write $Y = \mathbf{n}/a$.

Finally, to simplify the presentation of our analysis, we will always assume implicitly that continuous variables in our model are associated with well-behaved density functions – in particular, with continuous cumulative distributions.

3.2 Measures

We aim to define selection policies that lead to the selection of the candidates with the highest expected performance. At the same time, we aim to mitigate the disadvantage of certain demographic groups. Towards this end, we define the following measures.

Utility of Selection (UoS). A policy is useful if it leads to the selection of the candidates that actually perform best. This is captured by the following measure, defined in terms of the causal model.

$$\text{UoS} = \mathbb{E}[Y|T = 1]$$

Demographic Disparity (DmD). Our objective is to produce selection policies that minimize the difference between the selection rates of candidates of different backgrounds, according to their socio-economic features A . The following measure is defined for a group of candidates with sensitive properties $A = a$; it compares the selection rate for this group to the candidates outside the group.

$$\text{DmD}_A = P(T = 1|A = a) - P(T = 1|A \neq a)$$

The demographic disparity is equivalent to the *risk difference* measure from the literature [6], where risk is the probability of not being selected $1 - P(T = 1)$, and the difference is taken between the sensitive group and its complement.

3.3 Types of selection policies

To define alternative selection policies, we essentially have to define instances of the relationships $(A, X) \rightarrow S$ and $(A, S) \rightarrow T$, marked with thick arrows in Figure 1. In other words, for a given causal model, we aim to define the scoring function $S = f(A, X)$ and selection function $T = t(A, S)$ so that the policy exhibits desired values of the utility and fairness measures as described earlier. We say that every pair of functions $f(A, X)$ and $t(A, S)$ defines a *selection policy* within the *selection mechanism* defined by the model. We consider three types of policies: COEFFICIENTS, BONUS, and QUOTA.

COEFFICIENTS-based policies. This type of policy is completely determined by a series of coefficients w that are used to compute

the selection score S as a linear combination of the input scores \mathbf{X} . No sensitive attributes are used. The score is computed as

$$S = f(\mathbf{X} = \mathbf{x}) = \mathbf{w} \cdot \mathbf{x}, \quad |\mathbf{w}|_1 = 1. \quad (2)$$

Note that \mathbf{w} is normalized. Each component of \mathbf{w} represents the degree to which an input score counts towards the selection score. Candidates are admitted if their score exceeds a threshold τ , in which case we write $T = 1$ – otherwise, we have $T = 0$.

$$T = t(S) = \begin{cases} 1 & \text{if } S \geq \tau \\ 0 & \text{otherwise} \end{cases} \quad (3)$$

BONUS policies. This type of policy differs from **COEFFICIENTS** policies in that sensitive attributes \mathbf{A} contribute to the selection score of candidates. Specifically, the group of candidates with $\mathbf{A} = \mathbf{a}$ receive an additive bonus b in their score.

$$S = f(\mathbf{X} = \mathbf{x}, \mathbf{A}) = \begin{cases} \mathbf{w} \cdot \mathbf{x} + b, & \text{if } \mathbf{A} = \mathbf{a} \\ \mathbf{w} \cdot \mathbf{x}, & \text{if } \mathbf{A} \neq \mathbf{a} \end{cases}, \quad |\mathbf{w}|_1 = 1. \quad (4)$$

Note that the selection decision T is determined in exactly the same way as for **COEFFICIENTS** policies, i.e., via a threshold τ .

QUOTA policies. This type of policy determines the selection score in the same way as **COEFFICIENTS**, i.e., via a linear function with normalized weights \mathbf{w} . However, unlike **COEFFICIENTS** policies, sensitive attributes \mathbf{A} contribute to the decision T . **QUOTA** policies constraint the selected candidates from group $\mathbf{A} = \mathbf{a}$ to be (in expectation) a fraction $q \in [0, 1]$ of accepted candidates.

$$P(T = 1, \mathbf{A} = \mathbf{a}) = q \cdot \theta \quad (5)$$

Specifically, the selection decision T under a **QUOTA** policy is determined based on separate thresholds $\tau_{(\mathbf{A} \neq \mathbf{a})}$ and $\tau_{(\mathbf{A} = \mathbf{a})}$ for the two groups, $(\mathbf{A} = \mathbf{a})$ and $(\mathbf{A} \neq \mathbf{a})$, respectively.

$$T = t(S, \mathbf{A}) = \begin{cases} 1 & \text{if } (\mathbf{A} \neq \mathbf{a}) \text{ and } S \geq \tau_{(\mathbf{A} \neq \mathbf{a})} \\ 1 & \text{if } (\mathbf{A} = \mathbf{a}) \text{ and } S \geq \tau_{(\mathbf{A} = \mathbf{a})} \\ 0 & \text{otherwise} \end{cases} \quad (6)$$

The two thresholds are chosen so that Equation 5 is satisfied.

3.4 Problem definition

Having defined the set of policies that we consider, we can now define our objective formally. If we were to ignore the goal of favoring disadvantaged group $(\mathbf{A} = \mathbf{a})$ via an affirmative action policy, it would be natural to aim for the policy that optimizes the utility UoS of the selection mechanism. On the other extreme, if we were to target only the demographic disparity of those selected, it would be natural to aim for a policy that eliminates \mathbf{DmD} (i.e., leads to $\mathbf{DmD} = 0$). In general, we consider cases where we are willing to trade λ units of utility to decrease disparity by one unit. We thus define our technical objective as follows.

PROBLEM 1. For sensitive property \mathbf{A} , and two associated groups of candidates $(\mathbf{A} = \mathbf{a})$ and $(\mathbf{A} \neq \mathbf{a})$, define $f(\mathbf{A}, \mathbf{X})$ and $t(\mathbf{A}, S)$ so as to obtain a calibrated policy that maximizes

$$o(f, t) = \text{UoS}(f, t) - \lambda |\mathbf{DmD}_A(f, t)|, \quad \lambda \geq 0.$$

Notice that the objective function considers the absolute value of \mathbf{DmD} , as we wish to eliminate \mathbf{DmD} rather than maximize it in favor of the currently disadvantaged group.

3.5 Discussion on modeling choices

The causal model described above might not match exactly every real selection mechanism. We briefly discuss its assumptions in the context of university admissions.

Candidate decisions might not be independent. In our setting, the decision is made independently for each applicant based on their individual attributes. This modeling choice simplifies the analysis considerably, while it still allows us to understand the properties of the different policies. However, it does not hold in some cases. For example, for many universities there is a limit on the number of admitted students – therefore admissions are *not independent*, as one candidate’s admission can be another’s rejection.

Selection does not imply enrollment. We also assume that an applicant who is admitted by a university is bound to attend it. However, in reality applicants do not apply for admission to one institution only and might be admitted by several. In some systems, applicants are forced to prioritize their preferences for institutions and attend the first to admit them. In other systems, applicants can choose which one to attend. The model does not capture explicitly such admission systems. However, its assumptions work well for top universities that students are almost certain to attend if admitted. This is the case for our empirical analysis, which concerns a top university to which most selected applicants enroll (about 95%).

4 POLICY SEARCH FOR A SINGLE SENSITIVE ATTRIBUTE

This section describes algorithms to determine affirmative action policies for a single sensitive attribute \mathbf{A} . Section 5 extends them to multiple sensitive attributes.

4.1 COEFFICIENTS policy

Let us consider the case of $\lambda = 0$, in which the objective function simplifies to the measure of utility UoS . Let $m(\mathbf{x}) = \mathbb{E}[Y | \mathbf{X} = \mathbf{x}, T = 1]$ be the expected performance Y of a selected applicant with input scores \mathbf{x} . By straightforward calculations, we find that

$$\text{UoS} = \mathbb{E}[Y | T = 1] = \frac{1}{\theta} \int m(\mathbf{x}) \delta(f(\mathbf{x}) \geq \tau) d\mathbf{P}(\mathbf{x}). \quad (7)$$

Intuitively, Eq. 7 expresses that UoS is the average expected performance of applicants that exceed the threshold τ . Using $f(\mathbf{x}) = m(\mathbf{x})$ maximizes UoS over all possible scoring functions. In other words, if we could use the expected performance of an applicant as the scoring function, the resulting **COEFFICIENTS** policy would select those applicants that are expected to have the best performance Y and it would be optimal for $\lambda = 0$. However, we are constrained to use linear functions as scoring function $f(\mathbf{X})$, as defined in Equation 2.

In the case where $m(\mathbf{x})$ is indeed a linear function, then using $f(\mathbf{x}) = m(\mathbf{x})$ leads to an optimal **COEFFICIENTS** policy for $\lambda = 0$. To state the claim formally, let $F(S = s)$ be the cumulative probability distribution for selection S and $F^{-1}(\beta)$ the inverse cumulative distribution, i.e., the function that returns the score below which lies a fraction β of the population of applicants.

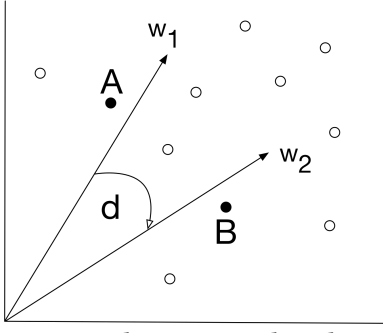


Figure 2: Given a weight vector w , the selection score of a candidate with input scores x is defined as the inner product $(w \cdot x) \propto \|x\| \cos(w, x)$. Each circle in the plot corresponds to the input scores of a candidate. Two candidates A and B have input scores with same l_2 -norm. A has higher selection score than B under score weight vector w_1 . However, under score weight vector w_2 , obtained by a rotation of angle d from w_1 , B has higher selection score.

LEMMA 1. Let $m(x) = \alpha_0 + \alpha c \cdot x$ with normalized coefficients, $|c| = 1$. Then, the following COEFFICIENTS policy is optimal for $\lambda = 0$:

$$\begin{aligned} S &= f(X, A) = c \cdot X \\ T &= t(S, A) = \begin{cases} 1 & \text{if } S \geq \tau \\ 0 & \text{otherwise} \end{cases} \\ \tau &= F^{-1}(1 - \theta). \end{aligned} \quad (8)$$

Conceptually, the optimal COEFFICIENTS policy for $\lambda = 0$ corresponds to ordering applicants in decreasing order of $c \cdot X$ and keeping the top θ -fraction of them by setting the threshold τ appropriately. By construction, this policy is calibrated.

For general $\lambda \geq 0$, c is not necessarily optimal anymore – instead, the optimal COEFFICIENTS policy would require different score weights. To see why, notice that, for two data points x_1, x_2 , we have

$$w \cdot x_1 > w \cdot x_2 \quad \text{iff} \quad \|x_1\| \cos(w, x_1) > \|x_2\| \cos(w, x_2).$$

Therefore, the selection score depends on the angle between input scores x and score weights w . In general, a rotation of score weights c by some angle d might lead to score weights w that change the ordering of datapoints $\{x\}$ by score value $\{f(x)\}$ – and thus possibly lead to smaller DmD and better objective value (see Figure 2).

To identify the optimal COEFFICIENTS policy we could use a sample of data $\{x\}$ and follow an approach similar to Asudeh et al. [2]. For a given set of applicants with input scores $\{x\}$, it is possible to partition the vector space of w into regions such that any vector of score weights w within the same region leads to the same ordering of applicants, where the ordering is performed by $\{w \cdot x\}$. However, we would have to perform an exhaustive search of such regions, which is inefficient for high-dimensional data. Instead, we experiment with a simple grid-like search strategy: starting with vector $w = c$, we rotate the score weight vector w by small increments towards different directions d , to obtain new weight vectors $w' = \rho(w; d)$ and evaluate the objective function in every step. The approach is described in Algorithm 1.

4.2 BONUS policy

Let M be the optimal COEFFICIENTS policy, having parameters c and τ as defined in Equations 8. We now consider a BONUS policy B that favors with b points the group $A = a$ that is *disadvantaged* under M , assuming without loss of generality that

$$P(T = 1 | A = a; M) \leq P(T = 1 | A \neq a; M).$$

For the moment, let us restrict B to use the same score weights c as M , but score threshold τ_b that is potentially different than the threshold τ used by M . We lift the restriction on score weights at the end of Section 4.2.

Calibrating τ_b . All policies are required to be calibrated, i.e., adhere to condition 1. Towards this end, let us consider a fixed bonus $b \geq 0$. The lemma below guarantees that there is an algorithm to choose its threshold τ_b so that B becomes calibrated.

LEMMA 2. The score threshold τ_b that makes B calibrated can be found via binary search over the interval $[\tau, \tau + b]$.

Optimizing b . Given τ_b obtained with the procedure described above, and $w = c$, how do we determine $b = b_{\text{OPT}}$ that optimizes the objective of Problem 1? Our main observation is that we can pre-determine a narrow range of values within which the optimal b_{OPT} must fall. To be precise, let $F(S = s | A = a; M)$ and $F(S = s | A \neq a; M)$ be the cumulative probability functions for the score S of the two groups under COEFFICIENTS policy M . Let $g(\beta) = F^{-1}(\beta | A = a; M)$ and $h(\beta) = F^{-1}(\beta | A \neq a; M)$ be the corresponding inverse functions: given a cumulative probability value β , they return the score S of the β -th applicant of the respective group. Then the optimal b_{OPT} lies in the interval given below.

$$\text{LEMMA 3.} \quad b_{\text{OPT}} \in I = [0, h(1 - \theta) - g(1 - \theta)]$$

PROOF. First, notice that the following condition holds for M .

$$P(T = 1 | A = a) \leq \theta \leq P(T = 1 | A \neq a) \quad (9)$$

Algorithm 1: Search for COEFFICIENTS policy

Parameters: $\{d\}$: set of rotation directions, k : number of rotations in each direction

```

1 /* initialize weights */
2  $w := c, c_{\text{opt}} := c$ 
3 for each direction  $d$  do
4   for  $i = 1..k$  do
5     /* rotate the vector of score weights */
6      $w := \rho(w; d)$ 
7     /* normalize the vector */
8      $w := w / \|w\|_1$ 
9     /* set the threshold  $\tau$  to make the policy calibrated */
10     $\tau = F^{-1}(1 - \theta)$ 
11    /* update best policy */
12    if  $o(w) > o(c_{\text{opt}})$  then
13       $c_{\text{opt}} := w$ 
14 return  $c_{\text{opt}}$ 
```

Algorithm 2: Grid-based search for b_{OPT} **Parameters:** k : granularity

```

1  $\epsilon = b_{DmD}/k$ 
2  $b_{OPT} := 0$ 
3 for  $i = 1..k$  do
4    $b = i \cdot \epsilon$ 
5    $\tau_i := \text{binarysearch}(\tau_{i-1}, \tau + b)$ 
6    $v := o(b, \tau_b)$ 
7   if  $v > v_{OPT}$  then
8      $b_{OPT} := b$ ;  $v_{OPT} := v$ 
9 return  $b_{OPT}$ 

```

This follows from the assumption that $A = a$ is the disadvantaged group and the fact that the overall proportion θ of accepted applicants must be between the proportions of accepted applicants of the two groups $A = a$ and $A \neq a$. It is easy to see that the proportion $P(T = 1 | A = a)$ of admitted candidates of the disadvantaged group is non-decreasing for increasing $b \geq 0$, while the corresponding proportion $P(T = 1 | A \neq a)$ for the advantaged group is non-increasing for increasing $b \geq 0$. This implies that, for increasing $b \geq 0$, there is a bonus value $b = b_{DmD}$ for which we have equality between the quantities of Condition 9 and $DmD = 0$. This is achieved for a bonus value that makes the $(1 - \theta)$ -th applicant of group $A = a$ have a score equal to the score of the $(1 - \theta)$ -th applicant of group $A \neq a$. Formally, we have $b_{DmD} = h(1 - \theta) - g(1 - \theta)$ and the BONUS policy with $b = b_{DmD}$ is calibrated for score threshold $\tau_{b_{DmD}} = F^{-1}(1 - \theta | A \neq a; M)$. The above implies that the absolute value $|DmD|$ of disparity decreases over the interval $b \in [0, b_{DmD}]$ and increases over $b \in [b_{DmD}, \infty]$. *Second*, it is easy to see that UoS is decreasing for increasing b . The two observations imply that the objective function (Problem 1) is optimized in $[0, b_{DmD}]$. \square

One simple algorithm to search for b_{OPT} is grid-based search, in which we evaluate the performance of BONUS (b) policies for uniformly distributed b in the range I . The calibration for each policy explored in this grid-based search can be done somewhat more efficiently based on the following generalization of Lemma 2, which allows us to narrow the range of binary search for the threshold τ that makes a policy calibrated. The search algorithm for b_{OPT} is shown in Algorithm 2 (proof omitted).

LEMMA 4. *For $i \in \{1, 2\}$, let τ_i be the threshold that makes the BONUS policy with b_i bonus points calibrated, and let τ be the threshold for M . If $b_1 \leq b_2$, then $\tau_1 \leq \tau_2 \leq \tau + b_2$.*

Solving Problem 1. So far in Section 4.2, we have constrained ourselves to BONUS policies that adopt fixed weights $w = c$. This raises the question: how do we choose w jointly with other parameters to find the optimal bonus policy for Problem 1? The answer is that we can simply use $w = c$, the optimal score weights for a COEFFICIENTS policy with $\lambda = 0$, as stated in following lemma.

LEMMA 5. *There is an optimal bonus policy with $w = c$.*

PROOF. *First*, consider two BONUS policies B_1 and B_2 with the following properties: (i) B_1 and B_2 are associated with the same value of $|DmD|$; (ii) B_1 uses score weights $w = c$, while B_2 does

not. Then, B_1 is associated with higher UoS and therefore higher objective value than B_2 . *Second*, notice that for well-behaved score distributions (specifically, for cumulative score distributions $F(S = s | A = a; M)$ and $F(S = s | A \neq a; M)$ that are continuous), it is possible to choose an appropriate bonus value $b \in (-\infty, +\infty)$ to achieve *any* value of $DmD \in (-1, 1)$, irrespective of score weights w . Combined, the above two arguments imply that for any BONUS policy B_2 with $w \neq c$, we can build a policy with same DmD and $w = c$, which is associated with higher objective value. \square

4.3 QUOTA policy

We find that there is a relationship of equivalence between BONUS and QUOTA policies, as pointed out in the following lemma.

LEMMA 6. *For a given BONUS policy, there is a QUOTA policy that leads to the selection of exactly the same candidates, and vice versa.*

PROOF. Let $P(T = 1, A = a)$ be the probability admitting applicants of group $A = a$ under a BONUS policy with bonus value b . Let also q_b be such that

$$P(T = 1, A = a) = q_b \theta$$

$$P(T = 1, A \neq a) = (1 - q_b) \theta.$$

and define the probabilities

$$\sigma_{A=a} = P(T = 1 | A = a) = \frac{q_b \theta}{P(A = a)}$$

$$\sigma_{A \neq a} = P(T = 1 | A \neq a) = \frac{(1 - q_b) \theta}{P(A \neq a)}.$$

We can construct a QUOTA policy with a quota $q = q_b$ for the disadvantaged group ($A = a$), by setting

$$\tau_{(A=a)} = F^{-1}(1 - \sigma_{A=a} | A = a)$$

$$\tau_{(A \neq a)} = F^{-1}(1 - \sigma_{A \neq a} | A \neq a),$$

as per the definition of QUOTA policies in Section 3.3. By construction, we have a QUOTA policy that accepts exactly the same applicants as the given BONUS policy. A similar argument can be made for the opposite direction. \square

This lemma allows us to focus our empirical study (Section 6) mainly on COEFFICIENTS and BONUS policies.

5 POLICY SEARCH FOR MULTIPLE SENSITIVE ATTRIBUTES

Our discussion so far has focused on policies that target one sensitive attribute. In this section, we introduce the problem of policy design for multiple sensitive attributes $\{A_i\}$.

PROBLEM 2. *For sensitive properties A_1, A_2, \dots, A_m , each partitioned into two associated groups of applicants ($A_i = a_i$) and ($A_i \neq a_i$), define $f(A_i, X)$ and $t(A, S)$ so as to maximize*

$$o(f, t) = \text{UoS}(f, t) - \sum \lambda_i |DmD_{A_i}(f, t)|, \quad \lambda_i \geq 0.$$

Algorithms. For COEFFICIENTS policies, Algorithm 1 is directly extensible to the setting of multiple sensitive attributes.

For BONUS policies, our goal is to assign bonuses b_1, b_2, \dots, b_m to the respective disadvantaged groups ($A_i = a_i$). One approach

is, similarly to Algorithm 2, to perform a grid search over possible bonus values for each sensitive attribute. However: (i) Lemma 2 does not extend directly to multiple dimensions, and so does not binary search; (ii) grid search at fine granularity and multiple dimensions is inefficient. Instead, we employ an incremental algorithm that works as follows: (i) at every step, the algorithm maintains bonus values b_1, b_2, \dots, b_m for each sensitive attribute, (ii) at every step, it greedily increments by an increment δb the bonus value b_i of the attribute A_i which leads to the best objective value.

For QUOTA policies, the equivalence with BONUS extends to multiple sensitive attributes (proof omitted), and so we do not explore them further here.

6 APPLICATION

In what follows, we present a large real-world dataset of university admissions (Section 6.1), build a causal model on it (Section 6.2), and evaluate the performance of affirmative action policies on the data (Section 6.3). In addition, we compare with two baseline approaches: (i) the FA*IR algorithm of Zehlike et al. [24], and (ii) the MEDIAN algorithm of Feldman et al. [9] and highlight interesting connections among them (Section 6.3).

6.1 Data

We analyze anonymized data about university admissions from Chile. In Chile, admission to undergraduate programmes is typically based on standardized tests, administered by a public entity, on language, mathematics, natural sciences, and human sciences, plus their grades in high school. The latter are converted to the same range of values as the standardized tests using a simple formula. Together, test results and grades form the input scores.

Universities announce their degrees (such as engineering, medicine, or law), and coefficients to convert the input scores into an admission score. For instance, the engineering school we analyze below uses 30% grades, 10% language, 45% mathematics, 15% natural sciences, and 0% human sciences. At the same university, the law school instead uses 40% grades, 25% language, 10% mathematics, 0% natural sciences, and 25% human sciences. After learning their test results, candidates rank degrees by decreasing order of preference. Then, applicants for each degree are ranked by admission score and the top ones are admitted.

We have access to two datasets from this admission process. The first corresponds to *national-level data* about students who took the standardized tests; the second corresponds to *university-level data* of students admitted to one of the most sought-after engineering degrees in one of the largest universities (95% of the admitted applicants enroll for this degree). The national-level data is provided by state authorities in anonymized form under a research agreement¹. This dataset includes the gender, high school type (public or private), income decile, grades in high school, and results in standardized tests of about 3,500 applicants who gave the test in 2017 and included the engineering degree of our university-level data among their preferences. The university-level data is similarly provided by university authorities in anonymized form specifically for research on discrimination and bias. For the university-level

¹ Department of Measurement, Evaluation, and Registry for Education (DEMRE). Data requests through <https://psu.demre.cl/portales/portal-bases-datos>.

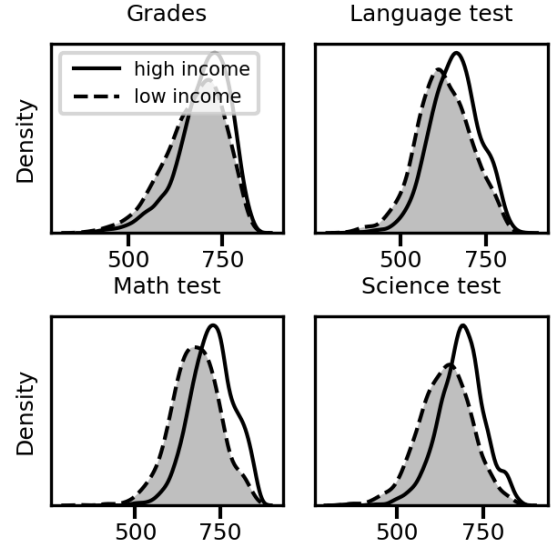


Figure 3: Differences in input scores by income.

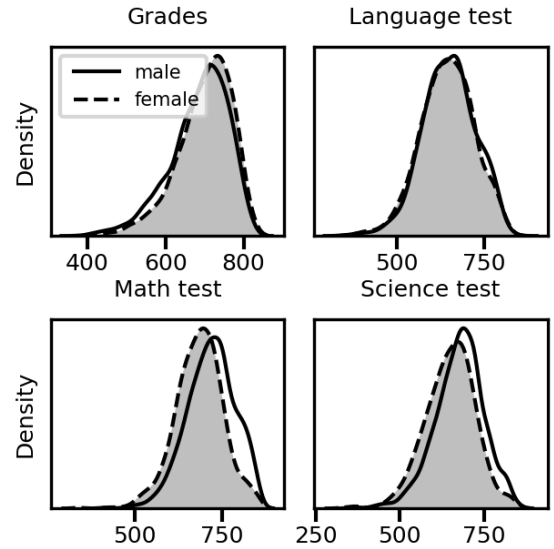


Figure 4: Differences in input scores by gender.

data, we have access to the same variables, except for income decile, for about 3,000 admitted applicants from 2010 to 2014. Most importantly, the university-level data contains the grades obtained by admitted students during their first year at the university. We note that the period of the two datasets do not overlap, but due to standardization of grades and test scores, the distributions of input scores do not differ significantly across years.

6.2 Building a causal model

The causal model (see Figure 1) is built from the two datasets we have at our disposal. In the causal model, there are two conditional relationships that need to be set to implement an admission policy,

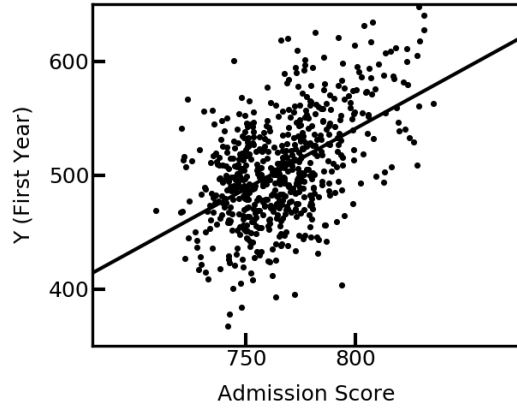


Figure 5: First year performance vs admission score of the optimal COEFFICIENTS policy with $\lambda = 0$. For clarity, the plot shows only a 20% sample of the data points, stratified by Y.

namely $(A, X) \rightarrow S$ and $(S, A) \rightarrow T$, and two conditional relationships that need to be learnt from data, namely,

$$A \rightarrow X : P(X | A) \quad (10)$$

$$X, T = 1 \rightarrow Y : P(Y | T = 1, X). \quad (11)$$

Both probabilistic quantities above can be extracted directly from the data (e.g., by fitting a statistical model on the involved attributes). Here is a subtle point that needs to be addressed for $P(Y | T = 1, X)$ (Expr. 11): if we learn a model for it from the data of students who were admitted ($T = 1$) according to a specific admission policy, can we use it to make predictions for students who *would be admitted* under a different policy – or are the observed data somehow *biased* by the decisions of the admission policy that generated the data? The answer is provided by causality theory [19], from which we deduce that we can, indeed, use the model we learn from the data to make predictions under different admission policies. The reason is that attribute X is an appropriate ‘control’ variable for Expr. 11 (more formally, X blocks all back-door paths from T to Y ; for an intuitive explanation see [19], Chapter 7).

In what follows, we use the attributes of gender (male or female), high school type (public or private), and income (above or below the median) as sensitive attributes A , with the population split in the two groups in parentheses; the values of grades, language test, math test, and natural sciences test as input scores X ; and the sum of grades during first year at university as Y (weighted by the credits of each subject; all first-year students take the same subjects).

For Expr. 10, we use directly the distributions extracted from the data, as shown in Figures 3-4 (the plots for school type are omitted here and whenever they are similar to those for income). As we see, the two groups for income (and similarly for school type) exhibit visibly different distributions (Fig. 3), while the distributions differ less for gender (Fig. 4). This suggests that, for the admission system that we study here, there is a stronger case for affirmative action policies that target financial attributes than gender, and the enactment of such policies has larger potential for impact.

For Expr. 11, we learn the best linear weights c for a COEFFICIENTS policy that targets only UoS (i.e., considers $\lambda = 0$) – they are

47.6% grades, 1.69% language, 23.8% mathematics, and 26.9% natural sciences. As discussed in Section 4.1, these weights also correspond to the linear function that expresses the expected performance Y of a selected applicant with given input scores. Figure 5 shows the actually observed values of admission scores S and performance Y of admitted students during their first year of studies. There is a correlation between the two, which supports our choice of a linear function to model the expectation $m(x) = \mathbb{E}[Y|X = x, T = 1]$ (see also beginning of Section 4.1).

Note also that there is variance in the observed values of Y for a given score value. Such variance does not affect the task of policy design in the setting we described in Section 3, as utility UoS is defined in terms of expected performance Y . Generally speaking, however, large variance in Y could trigger political arguments in favor or against affirmative action policies. On one hand, large variance might indicate that, instead of affirmative action, policy design should focus on the design of tests, the values X of which ought to be better predictors of future performance. On the other hand, variance that is too large compared to the expected performance of applicants might also support arguments in favor of affirmative action policies: if we cannot predict accurately who is going to do well at university, then we should not rely too much on test scores, but instead provide more opportunities to applicants from disadvantaged backgrounds.

Going back to our data, the distributions for admission score with weights c are shown in Figure 6. It becomes obvious that, for the COEFFICIENTS policy that targets only UoS, high-income applicants have an advantage over low-income applicants (similarly for private-school applicants against public-school applicants), in the sense that they are admitted by larger proportions. Male applicants have an advantage over female applicants, but the advantage is smaller than in the case of income or school type, consistently with our earlier observations.

6.3 Affirmative action policies

Figure 7 shows the empirical performance of different affirmative action policies, each policy concerning a single sensitive attribute. To produce the plot, we employed Algorithm 1 to search for alternative COEFFICIENTS policies and Algorithm 2 to search for different BONUS policies. The plots also show a few indicative quota percentages for the corresponding QUOTA policy of the same attribute. The plots also depict the optimal COEFFICIENTS policy with $\lambda = 0$, the best COEFFICIENTS and BONUS (and equivalent QUOTA) policy with $\lambda = 100$ discovered by the policy search algorithms 1-2, where λ is the parameter that dictates the trade-off between UoS and DmD in the optimization function (see Problem 1).

We make a few observations. (1) Based on its performance for income and school type, the BONUS and QUOTA policies provide better exploration of the trade-off between UoS and DmD: they lead to a better value of the optimization function (not shown in the plot) and policy instances with practically zero disparity DmD, with little loss in UoS. (2) The difference is less pronounced when gender is used as the sensitive attribute. This is explained by the fact that student grades show smaller discrepancy across genders than across income and school types (see Figure 3-4) (3) Consistently with the previous observation, the optimal BONUS policy uses larger

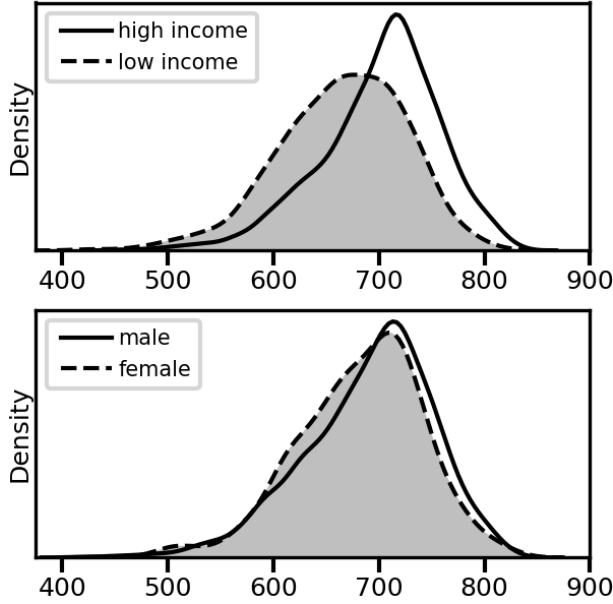


Figure 6: Distribution of admission scores for the optimal COEFFICIENTS policy with $\lambda = 0$.

bonus value for income and school type ($b = 30$ and 35 , respectively) than the one for gender ($b = 9.4$). To put these numbers in context, consider that the admission scores of the top applicant is about 833 points and of the last admitted applicant is in the range 722 - 733, depending on the attribute A . Hence, the gender bonus is negligible. (4) As expected, BONUS policies that lead to higher disparity (i.e., more favorable for the disadvantaged group) correspond to higher quota values for the corresponding QUOTA policy.

Baseline approaches We compare the policies discussed in this paper with policies based on alternative rankings produced by algorithms MEDIAN [9] and FA*IR [24].

MEDIAN [9] is applied on the scores X of each applicant and produces “repaired” feature values, such that it is impossible to distinguish from them whether an applicant belongs to group ($A = a$) or ($A \neq a$). Specifically, for each input score X , MEDIAN considers its distribution within each group, and produces repaired values such that an applicant of group ($A = a$) at the q -th quantile of the in-group distribution obtains the same repaired score as the corresponding applicant of group ($A \neq a$). This is done so that in-group order of applicants is preserved. For our purposes, we employ MEDIAN to obtain repaired scores X , use them to compute the admission score with weights c for all applicants, rank them and admit the top. In all cases, the result we obtain empirically is identical to the result of BONUS for which $DmD = 0$. This is explained theoretically: both policies admit top applicants from within each group – and by construction, the policy based on MEDIAN admits the same proportion of applicants from each group, leading to zero disparity.

FA*IR [24] takes as input the admission scores of applicants and produces a ranking that preserves the in-group rank of applicants, but at the same time satisfies “ranked group fairness” constraints. The latter ensures that within *any prefix* of the ranking both groups

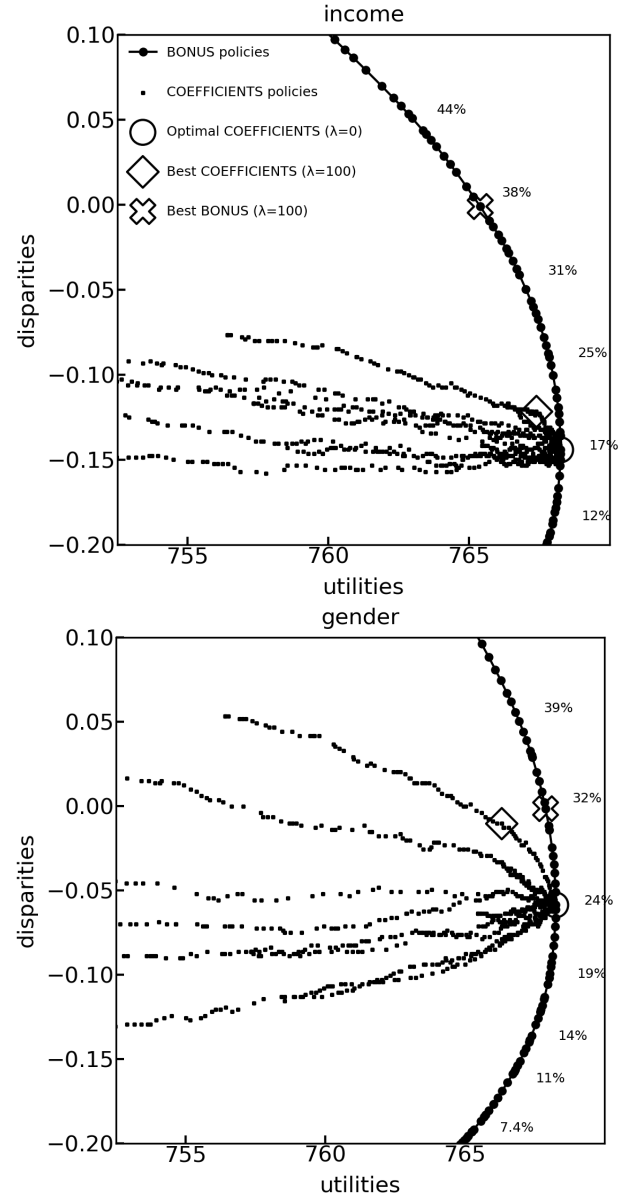


Figure 7: Comparison of COEFFICIENTS and BONUS policies. Quotas for the disadvantaged groups under QUOTA policies corresponding to BONUS policies are shown on the right.

are sufficiently represented – i.e., that the positions occupied by the disadvantaged group do not deviate significantly below a desired number. Deviation is defined in terms of a statistical test using two parameters α and p . For our purposes, we employed FA*IR on the admission scores obtained with weights c , and admit the top applicants. We observed empirically that both FA*IR and BONUS led to admissions that span the same utility-vs-disparity curve (see Figure 7). This is explained theoretically: both policies admit top applicants from within each group and, for appropriately chosen parameters, they can lead to the same desired number of admitted

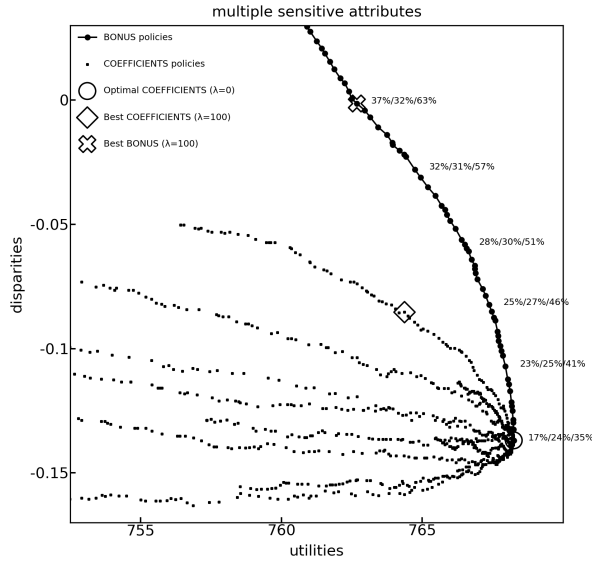


Figure 8: Performance of admission policies for multiple sensitive attributes. Percentages on the right are corresponding quotas for the disadvantaged groups (by income, gender, and school type) under Quota policies matching sample BONUS policies.

applicants from each group. However, we note that BONUS policy does not lead necessarily to the same *ranking* as FA*IR, even when it leads to the same *set* of k admitted applicants. This is because FA*IR enforces representation of the disadvantaged group for *every* prefix in the ranking, while BONUS ranks by selection score.

Multiple Sensitive Attributes Figure 8 shows the results when all three attributes (income, school type and gender) are considered. Again, we mark the optimal COEFFICIENTS policy for $\{\lambda_i = 0\}$, as well as the optimal COEFFICIENTS and BONUS (and equivalent Quota) policy with $\{\lambda_i = 100\}$, where $\{\lambda_i\}$ are the parameters that dictate the trade-off between UoS and DmD in the optimization function (see Problem 2). Consistently with the case of single sensitive attributes, the optimal bonus values assigned to income and school type ($b = 16$ and 26 , respectively) are larger than the one for gender ($b = 7$). We note that the baseline FA*IR cannot handle multiple sensitive attributes.

7 CONCLUSIONS

In this paper, we demonstrated how the problem of designing affirmative action policies can be addressed in an algorithmic framework. We explored three types of policies, i.e., COEFFICIENTS, BONUS, and Quota. We showed the equivalence of BONUS and Quota policies, and provided efficient search algorithms for the optimal parameters of COEFFICIENTS and BONUS policies. We analyzed real data for university admissions and obtained instances of admission policies that achieved good utility (performance of admitted applicants) while virtually eliminating disparity (i.e., the difference of the proportions of admitted applicants from advantaged and disadvantaged groups).

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